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Female Education, Marital Assortative Mating and Dowry: Theory and Evidence from India

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Abstract

This paper studies the association between female education, marital assortative mating and dowry in India. Literature suggests that marital assortative mating or ‘who marries whom’ is determined by the characteristics of grooms and brides and the respective households. Literature suggests that men choose to marry females with higher education in expectation of greater wage income in future. This paper proposes that in developing economies, characterized by low female labour force participation and poor returns to female education, female years of schooling is a non-monetary rather than a monetary trait. A better educated female brings down the household cost of production and is positively associated to household utility per unit of cost. Education levels of the potential bride on the other hand are negatively associated with dowry payments. A rise in female education would thus have countering effects on the utility of post marital household. A household would find positive assortative mating optimal if the efficiencies in household cost outweighs the fall in dowry with female education. Based on Becker (1991), we propose a basic theoretical model to explain empirical results. The empirical results for India using IHDS-II data suggest that there exists positive assortative mating in marriage based on education. The association however is found to be weaker in dowry prominent districts. There exists substitution in the education levels of bride and dowry. The cost reduction reinforcing effects are stronger than the dowry substitution effect. We derive instrumental variable estimates to address potential endogeneity.

Keywords: Female Education, Dowry, Marriage, Assortative Mating, Instrumental Variable Estimates

1 Introduction

Marriage is an inter-household social arrangement with an objective to maximise household utility. In a typical production function framework, household utility is a function of both monetary and non-monetary traits of the individuals and the households. Monetary traits such
as wages are associated positively to household income and utility. On the contrary, non-monetary traits in terms of years of schooling, physical attributes and intelligence determine efficacy in a household’s cost of production. These characteristics though do not contribute to household income and assets directly, but raise the ‘household output to cost ratio’ with a reduction in the total cost of production. The ability of husband and wife measured through education, intelligence, communication skills, etc. leads to time and effort reduction in production and generates synergies with increasing returns to scale. These factors besides monetary traits therefore gain substantial consideration at the time of marital arrangements, explicitly or implicitly. A household hence achieves greater utility with partners in the marriage market from a higher ability distribution.

In theory, household characteristics are also shown to be a strong predictor of marital assortative mating. Literature provides evidence of a strong positive association between household assets, age of the couples, height, etc. Rosenzweig and Stark (1989) show that household wealth and variation in risks to income determine the quality of marital matches. Deolalikar and Rao (1992) in empirical estimates from south-central India find a positive association between individual and household characteristics in marital sorting.

This paper studies marital assortative mating based on female years of schooling in developing countries with reference to Becker’s (1991) model on theory of marriage. Literature on assortative mating in education argues that men choose to marry females with higher education in expectation of increased household income. However, in developing countries such as India, where female labour force participation is as low as 27 percent, with poor female literacy and low returns to female human capital investment, an explanation for positive assortative mating based on education (found in empirical studies for India) lies beyond the wage market. This means that men with higher education see benefits other than future wage income from marrying females with greater years of schooling. Assortative mating in education may also reflect non-monetary returns to human capital. Potential brides

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1 Marital assortative mating is defined as ‘who marries whom’. Marital choices are based on individual and household characteristics such as education, income, physical appearance, etc. When individuals choose to marry from their own characteristic distribution, for example, a tall individual finds it optimal to marry someone who is also tall, is known as positive assortative mating. On the contrary, if a male with high wage finds it optimal to marry a female with lower wages, such that the male has a competitive advantage in labour market and female in managing household chores and children, this is known as negative assortative mating. Different individuals may find positive or negative assortative mating on a given trait optimal, depending upon how the trait enters into their utility function.

2 Census of India, 2011.
with higher education bring down the cost of household production and contribute positively
to household production, per unit of cost. Low education level of potential bride adds
positively to household utility as a non-monetary rather than a monetary trait.

Besides individual characteristics, in countries such as India, marital choices are also
determined by strong social and cultural norms. Marital decisions are mostly a family
phenomenon with strong parental controls and involve marital transfers such as ‘dowry’, an
important determinant of post-marital household utility levels. To the best of our knowledge,
the literature so far does not assess the impact of female education on marital sorting with
changes in relative bargaining powers of potential grooms and brides reflected in average
dowry exchanges. In India, characterised by patriarchal norms, groom price or dowry exists
as a social culture. It is a transfer made by the bride’s family to the groom’s family in cash or
kind, or the excess marital expenditure borne by the bride’s family relative to the groom. In
economic theory, dowry in the marriage market is perceived as an equilibrating mechanism to
compensate for the differences in the characteristics of the bride and the groom (Anderson,
2007). Any changes in the education level of females would alter the ability gap between the
potential groom and the bride. This would affect the marital transfer equilibrium levels and
assortative mating decisions. Marital assortative mating outcome of female education would
thus be a net effect of positive marginal utility with reduction in cost of household production
and fall in marginal utility with fall in dowry.

There are four main aims of this paper. First, we study the presence of assortative mating by
looking at the association between years of schooling of husband and wife in India, using
data from India Human Development Survey II (IHDS-II). This is an important study for a
country such as India with historical patriarchal norms and female gender biases, to see how
female human capital investment is perceived in the marriage market. A positive association
would indicate societal preferences for bride education which would have further
implications for socio-economic outcomes such as intra-household asset sharing, gender
parity and household decisions³. On the other hand, if the association is negative, it would
indicate absence of net positive utility from bride education levels in marital ties and a
disincentive to invest in female education. Second, we study if assortative mating on
education differs for districts with and without prominence of dowry. This may be true
because dowry exists as a social norm in India and dowry may be correlated with assortative

³ Study of socio-economic outcomes of marital assortative mating is not in the scope of this paper.
mating choices. Districts with higher prominence of dowry may view bride education and dowry income as substitutes, such that high amounts of dowry could substitute for lower bride’s schooling years. Marital assortative mating outcomes should be weaker in such districts. Third, we study the complementarities or substitutability between spousal education levels and their association with dowry expenditure. If an educated bride’s family is willing to pay a higher dowry for a higher educated son-in-law, there would exist complementarities in the bride and groom education levels, and would have a positive association with dowry transfers. However, if the bride’s family is less willing to pay a dowry because the daughter is educated or the groom’s family is willing to accept a lower dowry for an educated bride, husband and wife schooling years are substitutes in dowry. In this case dowry transfers would fall with incremental education years of husband and wife together. Fourth, we propose a basic theoretical model to explain the scenarios under which positive or negative assortative mating would be an ideal outcome. The conditions are derived for the post-marital household of the groom’s family with an objective to maximize household utility in presence of dowry.

There are several papers that study the association between husband and wife’s level of education. Most of the literature argues that assortative mating based on education is determined by the labour market returns and monetary yields of human capital investment. Pencavel (1998) states that assortative mating by schooling strengthened since 1960. The paper attributes the increase in similarity of schooling levels mainly to increased work opportunities for females. Greenwood et al. (2014) find similar results where Kendall’s $\tau$ rank correlation$^4$ between the education of husband and wife is increasing between 1960 and 2005.

Average monetary returns to education however may not completely account for assortative mating patterns in developing nations. Female education is a human capital endowment with substantial non-monetary benefits besides labour market returns in terms of raising children, their education, health and sanitation, support to household, financial and social decisions, etc. Mother’s schooling has been shown to have substantial positive effects on children’s schooling. The effects have in fact been found to be larger than father’s schooling (Qian, 2008; Behrman, 1997).

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$^4$ Kendall’s Tau is a non-parametric rank correlation based on the similarity of ordering of data ranked on each of the quantities. Calculations are based on concordant and discordant pairs.
Theoretically, assortative mating decisions could be analysed in the framework provided by Becker (1991). The model states that ‘who marries who’ is a function of joint marginal effect of traits of potential grooms and brides and their respective households. If the traits are complementary which means they offer increasing returns to scale in the household production, a household would find positive assortative mating to be desirable. Negative assortative mating in the marriage market may also be an ideal strategy if the traits of potential groom and bride are substitutes to each other.

This paper suggests that education as a trait for potential brides and grooms may act as a substitute if years of schooling is a marketable characteristic, with similar returns in education to both male and female. In this case, both husband and wife may want to allocate more time in the labour market vis-à-vis household chores. This may generate substitution effects in the labour market. This argument gains support from Gihleb and Lang’s (2016) study for the US where the estimates show that the educational homogamy has not increased in the country, thereby suggesting that male and female do not find it optimal to marry in similar educational distribution due to the presence of potential substitutability in labour work hours. In countries such as India, where market returns to female education are substantially poor and female workforce participation is as low as 27 percent and even lower for married women, i.e., 20 percent (Census of India, 2011), there exists a strong case for a nearly perfect segregation of roles between husband and wife in market and household work, respectively.

Societies based on patriarchal norms allocate higher weightage to female household roles, rearing of children and attach social stigma to female wage labour, such that husband and wife’s education is viewed as a complement. If the husbands have a comparative advantage in labour market and wives in managing household and children, years of schooling generate complementarities in marital ties in such socio-economic environment. Boulier and Rosenzweig (1984) provide evidence for Philippines where with low female labour force participation, there exist positive payoffs for positive assortative mating in the presence of education as a human capital endowment with non-monetary returns. In contrast, Dalmia and Lawrence (2001) find that age was a stronger predictor of assortative mating in India, and education in the US. However, Anukriti and Dasgupta (2017) in a review of literature on

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5 Positive assortative mating is where individuals choose spouses in the marriage market from an ability distribution similar to their own. For example, higher educated marry those with high levels of education, taller individuals choose tall spouses, etc.

6 IHDS-II ‘Eligible Women’ survey data suggests that about 40% of women are not willing to work even if found a suitable job. And nearly 40% of females denied being allowed to work if found a suitable job.
economics of marriage market outcomes in developing countries observe strong homogamy in education. Empirical support on positive assortative mating in education in countries with very low female workforce participation indicates the presence of non-monetary benefits of education for households.

This paper further recognises the association between female education and dowry, and its impact on assortative mating. Dowry may amount to a significant proportion of household wealth as a one-time transfer. Such inter-household transfers act as income insurance in low income countries (Rosenzweig and Stark, 1989). Anderson (2003) shows that with an increase in the average income and modernisation, there is a decline in average dowry payments. However, in India, where there exists caste hierarchy and diversification in the share of social and economic power, modernisation and rise in income is shown to have led to an increase in dowry payments. Rao (1993b) in estimating the rising groom prices in South-Central India shows that dowry transfers rise with the widening gap in bride and groom characteristics such as age, height, land ownership and years of schooling. However, Edlund (2000) in a comment to Rao (1993b) argues that individual traits provide a better fit to dowry estimation as compared to the difference in traits. A positive shift in the potential brides’ ability distribution and rise in female education is expected to raise female bargaining power in the marriage market. Sweeney and Cancian (2004) find significant results of a shift in marriage market equilibrium in favour of women with increased earning potential in the US. Therefore, a relative change in the characteristics of the potential brides and grooms either through education, workforce participation or other traits is believed to alter marriage market outcomes in terms of assortative mating and dowry. Anderson’s (2007) model derives dowry as an equilibrating factor in the marriage market for gaps in characteristics of the potential groom and bride. The larger is the gap in desirable traits, the higher is the dowry payment.

According to this paper, dowry may compensate for lower education levels of the potential bride. In the presence of low labour market returns to female education and substitutability between dowry and years of schooling, negative assortative mating in years of schooling may be an ideal outcome for the households. However, in the presence of non-monetary human capital endowments that generate positive marginal utilities in household production, positive assortative mating may be optimal if this effect is stronger than the substitution effect between female education and dowry. There exists a trade-off between dowry and household cost reduction with rise in female education. This paper provides a basic theoretic model for
marital decisions on female education in the presence of countering effects through dowry and efficiency in household costs. Empirical support on positive assortative mating suggests that positive marginal productivity of female education outweighs the negative effects of dowry substitution.

Given a comparative advantage in female education as a non-monetary rather than a monetary trait in developing nations, we test the proposition that brides and grooms with higher schooling marry with potential mates from a similar ability distribution. Marginal productivity of female education is higher than the marginal fall in dowry. The results may however be weak or in contradiction for districts with greater prominence of dowry. The substitutability between female education and dowry may be stronger where dowry persists as a strong social norm. We derive empirical estimates from IHDS-II 2011 data, a household level survey for the districts of India, and the results lend support to our hypothesis. Empirical findings are supported through a basic model of ideal marital sorting decisions with effects through dowry.

Following the introduction, rest of the paper is organised as follows: theory in section II; data and empirical strategy in sections III and IV respectively; empirical results and validation tests in section V; and conclusive comments in section VI.

2 Theoretical Model

This section provides a basic theoretical model to explain marital assortative mating based on education and the role dowry. This model is based on Becker’s *Assortative Mating in Marriage* (1991) model which suggests that a household maximizes production per unit of cost subject to the budget constraint. Our model derives conditions under which positive or negative assortative mating would be optimal based on education and in the presence of marital transfers, such as dowry. In this case, groom’s household is the decision-making household that maximizes production and determines in relation to the groom’s education, whether to seek a bride with a higher education (positive assortative mating) or a bride with lesser education (negative assortative mating). Driven by the data limitations, household production function of the bride’s family is not considered in this model.

The groom’s household utility function is defined as a function of $m$ market goods and services $x_i$, and leisure $t_j$ of each male individual $j$.

*Household Production Function*
\[ Z = f(x_1, \ldots, x_m; t_1, \ldots, t_j) \]  

where,

\[ x_i: \text{various Market Goods and Services; } \quad i = 1 \ldots m \]

The household is constrained on budget such that the total expenditure on consumption goods and services, \( x_i \), with price vector \( p \) cannot exceed the sum of income from labour \( l_j \) at a wage rate \( w_j \) for each male individual \( j \) in the household\(^7\), property income \( v \) (given and constant) and dowry \( D \). Dowry is a one-time inter-household transfer and adds to the income of the post marital household (groom’s household). Dowry is a function of male and female education in the post-marital household, \( S_m \) and \( S_f \) respectively.

**Budget Constraint**

\[
\sum_{i=1}^{m} p_i x_i = \sum_{j=1}^{n} w_j l_j + v + D(S_m, S_f); \tag{2}
\]

where,

\[ p_i: \text{price of market goods and services} \]

\[ w_j: \text{wages of male individual } j \]

\[ l_j: \text{work hours of male individual } j \]

\[ v: \text{property income} \]

D: dowry

\( S_f: \text{Education of female, Bride} \)

\( S_m: \text{Education of male, Groom} \)

**Time Constraint**

\[ l_j + t_j = T, \quad \forall j \tag{3} \]

Substituting (3) in (2), the goods market and time constraints can be transformed into full income constraint as:

\[
\sum_{i=1}^{m} p_i x_i + \sum w_j t_j = \sum w_j T + v + D(S_m, S_f) = I \tag{4}
\]

\(^7\)The model assumes that only males participate in the wage market.
where \( I \) stands for full income achievable with \( w_j \) as given in a perfectly competitive labour market.

In developing nations characterised by low female workforce participation and poor returns to female education, there exists division of work roles between males and females. Males take up larger roles in the labour market and females in household activities. The model therefore assumes that men participate in the labour market and females’ education and other traits are non-monetary in nature. It is also assumed that wages are competitive across labour markets such that, \( w_m = w \).

**Household Cost of Production**

Household production necessitates the employment of inputs in the production process. The household output is the amount of goods and services it could consume, assets it could raise and the quality of children it could rear\(^8\). Costs would involve time, effort and human capital required in the process of production. Cost of household production is a function of prices of consumption goods and education of household members. Rise in prices makes consumption expensive, raising the opportunity cost of leisure and cost of household production. Education enhances human efficiency through increased marginal productivity in production outcomes of health, child education, etc., leading to lower costs of household production.

\[
C = C(p, S_m, S_f) \tag{5}
\]

With improved efficiency in education, marginal cost of the household is assumed to be decreasing in the education of both male and female such that:

\[
\frac{\partial C}{\partial S_m} \equiv C_m < 0
\]

\[
\frac{\partial C}{\partial S_f} \equiv C_f < 0
\]

The household production could be represented as a proportion of household production to costs.

\[
Z = \frac{\text{Full Income}}{\text{Average Cost of Production}} = \frac{\sum w_j T + v + D(S_m, S_f)}{C(p, S_m, S_f)}
\]

\(^8\) Children could be considered as consumption goods that generate positive utility with inputs required in terms of time and effort.
The greater is the production per unit of cost, the higher is the return and the utility that a household could achieve. As is evident from the household income per unit of cost, the education of the potential grooms \((S_m)\) and brides \((S_f)\) impacts household production function both via a change in dowry income and a fall in the cost of household production. Education therefore enters the model as a non-monetary trait, such that it affects the household output to cost ratio only via reduction in the household cost of production and dowry, a one-time marital transfer. A rise in male education raises the groom-price in the market with a corresponding increase in dowry transfers. Higher male education simultaneously brings down the cost of household production. The total effect of increased male education is thus positive on household production per unit of cost. Female’s education on the other hand with increased bargaining power in the marriage market bears opposing effects on the production to cost ratio. The groom household would witness a fall in dowry with rise in female education, reducing the post-marital income and a fall in cost with positive effects on per unit returns. The net effect of female education on household production per unit of cost is ambiguous. Mathematical explanations for the same are provided in the following sub-sections.

**Impact of Education on Total Output**

To study the effect of education levels of bride and groom on household production per unit of cost, we derive the first order partial derivatives. It is to be noted, that education level of the bride is her education status at the time of marriage. We assume that the household decision on assortative mating based on education is not on investing in the education of girl after marriage, but on choice of a female with a given level of education among potential brides (with different education levels) in the marriage market.\(^9\)

**First Order Partial Derivatives**

With the assumption of constant wages in the labour market, education impacts total achievable income \(I\), only through the effect on dowry.

\[^9\text{We run a robustness test for empirical findings on assortative mating to eliminate any possibility of investment in bride education after marriage.}\]
\[
\frac{\partial I}{\partial S_m} = \frac{\partial D}{\partial S_m} = D_m
\]
\[
\frac{\partial I}{\partial S_f} = \frac{\partial D}{\partial S_f} = D_f
\]

A marginal increase in groom’s education leads to an unambiguous increase in the production function.

\[
\frac{\partial Z}{\partial S_m} = \frac{\partial [D(S_m, S_f)C^{-1}]}{\partial S_m}
\]
\[
= \frac{\partial D}{\partial S_m} C^{-1} - C^{-2} D(S_m, S_f) \frac{\partial C}{\partial S_m}
\]

\[D_m - C^{-1} C_m D(S_m, S_f) > 0 \quad (6)\]

The result from equation 6 would hold with the underlying assumptions of rise in dowry and fall in cost with male education respectively. This implies \(D_m > 0\) and \(C_m < 0\). Further, dowry that is the net transfer from bride’s family to groom’s family is positive such that \(D(S_m, S_f) > 0\). A rise in male education therefore causes a definite positive effect on household income. Therefore, investment in human capital of sons is found to be profitable by the households.

This could also be written in terms of elasticities:

\[\tau_m - \omega_m > 0 \quad (7)\]

(Derivation in Appendix A.1)

where \(\tau_m = \frac{\partial D}{\partial S_m} \cdot \frac{S_m}{D}\) and \(\omega_m = \frac{\partial C}{\partial S_m} \cdot \frac{S_m}{C}\)

Therefore, the total output would rise with male education if the male education elasticity of dowry and household cost are positive together (which would always be the case since \(\omega_m < 0\)). An increase in human capital investment by the male counterpart unambiguously leads to an increase in household production, both through an increase in dowry that adds to the income of the household and a decrease in the cost of household production.

Similarly, the marginal effect of female’s education would result in an increase in total household output if:
\[ D_f - C^{-1} C_f D(S_m, S_f) > 0 \]

i.e.

\[ \tau_f - \omega_f > 0 \] (8)

By assumption, since \( \tau_f < 0 \), thus the total output of the household would increase with female education, if the proportionate fall in household costs is greater than the proportionate fall in dowry.

\[ |\omega_f| > |\tau_f| \] (9)

The independent effects of male and female education on household provide easy interpretations. The model shows that a household always gains from investing in male’s education since it is positive in both dowry and cost reduction. A female’s education on the other hand leads to countering effects by negatively impacting dowry and positively on cost reduction. Thus, there exists a trade-off in decisions on the education level of potential brides. A household would be more acceptable of a bride with higher education if the efficiency gains in cost outweigh the loss in dowry component.

**Assortative mating**

Mate seeking in the marriage market is a function of different characteristics of the potential grooms and brides and their respective households.\(^{10}\) On one hand where the household production is monotonic and increasing in education, assortative mating, i.e., who marries whom amounts to joint marginal effects on household production. As Becker (1991) explains in the paper, if increasing both \( S_m \) and \( S_f \) increases the total output, \( Z \), by the same amount, as the sum of addition when each is increased separately (in the presence of constant returns to scale, CRS), all sorting of male and female would give same total output, \( Z \). However, with increasing returns to scale (IRS), sorting of large \( S_m \) with large \( S_f \) and small \( S_m \) with small \( S_f \) would give greatest total output. Mathematically, positive assortative mating in education, mating of likes (highly educated marry the highly educated and those with lower

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\(^{10}\) Given the objective of the model, it is assumed that the characteristics of the households such as wealth, location, etc. are similar across the households.
levels of education marry the ones with lower education) is desirable if \( \frac{\partial^2 Z(s_m,s_f)}{\partial s_m \partial s_f} > 0 \). And negative assortative mating of un-like is desired if \( \frac{\partial^2 Z(s_m,s_f)}{\partial s_m \partial s_f} < 0 \).

A household would gain from positive assortative mating in education of male and female if the traits act as complements. Those with higher education would seek better educated spouses if the efficiencies achieved with higher education in cost outweigh the fall in dowry with higher female education. Dowry may also lead to an increase in the education level of both male and female if the parents of the bride are willing to pay a higher dowry to secure a groom with higher education than the daughter’s. It would not be ideal for a groom with lower education to seek a better educated bride if the cost reduction is not sufficient to compensate for the fall in dowry.

Dowry compensation by the bride’s family may offset the cost inefficiencies associated with lower education levels of the girl. In such scenario, households find negative assortative mating as the desirable strategy. The marriage market would thus witness larger gaps in the education levels of husband and wife.

The cross partial effects of male and female education would determine whether marriage sorting would be positive where males and females seek potential partners from similar ability distribution or different. Positive marriage sorting would require the cross partial derivative of \( Z \) to be positive, such that the total output increases with the education of both the male and the female in the household. Negative sorting would be an ideal decision for a household if an increase in female’s education with a rise in male’s education is causing the total output to fall such that the education traits are substitutes.

**Second Order Cross Partial Derivatives**

\[
\frac{\partial^2 Z}{\partial s_m \partial s_f} = C^{-1} D_{m,f} - C^{-2} D_m C_f - C^{-2} D_m D_f + 2 C^{-3} C_m C_f D - C^{-2} C_{m,f} D \tag{10}
\]

(Derivation in Appendix, section A.2)

Where \( D_{m,f} = \frac{\partial D}{\partial s_m} \) and

\footnote{Owing to the cultural norms in developing countries such as India, it is preferred to marry a boy with higher traits such as education, income, etc. as compared to the bride.}
\[ C_{m, f} = \frac{\partial \left( \frac{\partial C}{\partial S_m} \right)}{\partial S_f} \]

Since \( C^{-1} > 0, D_m > 0, D_f < 0, C_m < 0 \) and \( C_f < 0 \), thus the effect of female education on marriage sorting would depend upon the relative magnitude and signs of \( D_{m,f} \) and \( C_{m,f} \).

Education may have either reinforcing or offsetting effects on household cost of production and magnitude of dowry. The cost effects are reinforcing if the education of male and female act as complements, where female education leads to a greater reduction in marginal costs of male education such that \( C_{m,f} < 0 \). However, such costs could be offsetting if the household costs of production fall at a decreasing rate, such that \( C_{m,f} > 0 \). For example, if a male is not competitive at managing household chores or children’s education, marrying an educated female with greater productivity in such activities would bring down the household cost of production at a greater rate. However, if education makes a male better in managing household chores or children’s education, then an educated female would act as a substitute. The costs though would fall due to the individual effect of female’s education, but the rate of fall would be lesser.

Under the assumption of dowry as an equilibrating factor in the marriage market, the magnitude of dowry should fall with the rise in female education, i.e., \( D_f < 0 \). The marginal effect on dowry however with changes in both male and female education is less obvious. If dowry and female education are substitutes in the marriage market, an increase in both male and female education would cause a marginal decrease in dowry. Education would have offsetting effects such that \( D_{m,f} < 0 \). However, if the two are viewed as complements, the market would witness an increase in dowry with the education level of the groom and the bride. This could still hold under the assumption that \( D_f < 0 \). Though on an average, dowry may fall with the education level of the female, yet an educated female may be willing to pay a higher dowry for a better educated male. This may be supported by increased ability to pay a higher dowry through enhanced female education level and workforce participation. Female education may therefore generate complementarities in the marriage market for dowry such that \( D_{m,f} > 0 \).

The appropriateness of positive and negative sorting in the marriage market would determine whether the household costs and dowry are offsetting or reinforcing the education trait of potential grooms and brides.
Case I: Offsetting household cost and dowry in education of male and female, $C_{m,f} > 0$ and $D_{m,f} < 0$

Equation 10 suggests that households in the marriage market would find positive sorting to be desirable, i.e.,

$$\frac{\partial^2 Z}{\partial S_m \partial S_f} > 0,$$

if

$$2C^{-2} C_m C_f D - C^{-1} D_m C_f > C^{-1} C_{m,f} D + C^{-1} C_m D_f - D_{m,f}$$

(11)

This could also be written as

$$2 \omega_m \omega_f - \tau_m \omega_f > \omega_{f,m} \omega_f + \omega_m \tau_f - \tau_{m,f} \tau_m$$

(11’)

where $\tau_m = \frac{\partial D_m}{\partial S_m} \frac{S_m}{D}$; $\omega_m = \frac{\partial C_m}{\partial S_m} \frac{S_m}{C}$; $\omega_{f,m} = \frac{\partial C_f}{\partial S_m} \frac{S_m}{C_f}$; $\tau_{m,f} = \frac{\partial D_m}{\partial S_f} \frac{S_f}{D_m}$

(Derivation in Appendix, section A.3)

Under the assumptions of falling marginal costs in education ($C_f < 0, C_m < 0$) and rise in dowry with male education ($D_m > 0$), the left side of equation 11 is always positive. If the costs are offsetting such that $C_{m,f} > 0$, the marriage market can still achieve positive sorting where the likes get sorted in equilibrium if the multiples of direct effects of fall in cost and rise in dowry with education are stronger than the offsetting effects on cost and dowry. This would be true if the individual efficiencies in cost achieved with higher education are stronger than the offsetting effects on both cost and dowry. If the bride individually is able to bring down the household costs to an extent that it compensates both for the loss in dowry and any substitutions in the cost reductions from male education, positive assortative mating would be ideal even with offsetting education effects on costs and dowry. If direct elasticities of cost are greater than cross elasticities of cost and cross elasticities of dowry in male and female education, then positive assortative mating would be desirable for a household.

Case II: Offsetting household costs and reinforcing dowry effects on education of male and female, $C_{m,f} > 0$ and $D_{m,f} > 0$

With $C_{m,f} > 0$ and $D_{m,f} > 0$, households in the marriage market would find positive sorting to be ideal, i.e.,
\[ \frac{\partial^2 Z}{\partial S_m \partial S_f} > 0, \text{ if} \]

\[ 2C^{-2}C_mC_fD - C^{-1}D_mC_f + D_m > C^{-1}C_{m,f}D + C^{-1}C_mD_f \]  \hspace{1cm} (12)

This could also be written as

\[ 2\omega_m\omega_f - \tau_m\omega_f + \tau_{m,f} > \omega_f\omega_f + \omega_m\tau_f \]

The result of positive assortative mating gets strengthened when the dowry effect is reinforcing. Dowry effect may be reinforcing, i.e., dowry increases with the education of male and female together, with the willingness to marry daughter to a higher educated groom. Increased female education levels may also translate into higher ability to make dowry payments with increased probability of workforce participation. The household would achieve higher output with an increase in education levels of both the groom and the bride if not only are the individual cost reducing effects stronger than the offsetting effects on cost, but dowry increases with the education of male and female together.

**Case III: Reinforcing household costs and offsetting dowry effects on education of male and female, \( C_{m,f} < 0 \) and \( D_{m,f} < 0 \)**

With \( C_{m,f} < 0 \) and \( D_{m,f} < 0 \), households in the marriage market would find positive sorting to be desirable, i.e.,

\[ \frac{\partial^2 Z}{\partial S_m \partial S_f} > 0, \text{ if} \]

\[ 2C^{-2}C_mC_fD - C^{-1}D_mC_f - C^{-1}C_{m,f}D > C^{-1}C_mD_f - D_m \]  \hspace{1cm} (13)

This could also be written as

\[ 2\omega_m\omega_f - \tau_m\omega_f - \omega_f\omega_f > \omega_m\tau_f - \tau_{m,f}\tau_m \]

It is possible that an increase in male and female education may have reinforcing effects on the cost of production where the education traits act as complements. However, with a rise in female education, not only does the dowry component fall, but witnesses an offsetting effect. This is generally true as females begin to achieve a higher socio-economic status such that males, specifically the more educated, are willing to accept lower dowry for a girl with higher education. In such a scenario, for positive assortative mating to be a desirable outcome, it is
required that the fall in costs—both direct and cross—should be strong enough to overcome the fall in dowry with a rise in female education and the corresponding offsetting effects.

Case IV: Reinforcing household costs and dowry effects on education of male and female, $C_{m,f} < 0$ and $D_{m,f} > 0$

With $C_{m,f} < 0$ and $D_{m,f} > 0$, the households in the marriage market would find positive sorting to be ideal, i.e.,

$$\frac{\partial^2 Z}{\partial s_m \partial s_f} > 0,$$

if

$$2C^{-2}C_mC_fD - C^{-1}D_mC_f - C^{-1}C_{m,f}D + D_{m,f} > C^{-1}C_mD_f$$  \hspace{1cm} (14)

This could also be written as

$$2\omega_m\omega_f - \tau_m\omega_f - \omega_{m,f}\omega_f + \tau_{m,f}\tau_m > \omega_m\tau_f$$

In a scenario where both the household costs and dowry are reinforcing due to male and female education, positive sorting would have a stronger likelihood. In a country such as India, dowry is a cultural norm. Though it may vary and be determined by various factors, with the patriarchal structure and poor socio-economic status of females, a positive groom price is a social standard. On one hand where a rise in female education may lower the dowry component on average, the parents of the bride may be willing to pay a higher dowry for a better, higher educated groom. Bringing a bride with higher education in the family would thus add to the household production, both by reduction in cost and a rise in dowry.

The theoretic model of marital assortative mating and dowry implies that:

a. Education contributes positively to household utility as a non-monetary trait.

b. There exists a trade-off and potential substitutability between dowry and female education.

c. In patriarchal societies where dowry exists as a cultural norm, marital choices are governed by differences in marginal utilities to education and dowry.

3 Data

With an objective to examine marriage market sorting and marital decisions on the education level of potential grooms and brides in India, this paper uses household level data from IHDS-II and Census of India. We use data from IHDS-II data available for the year 2011.
IHDS-II is a nationally representative, multi-topic survey of 42,152 households across India for 2011-12\textsuperscript{12}. The survey gathers extensive details on variables related to health, education, employment status, marriage and other components that are important indicators of economic and social well-being. This paper in particular maps data at the level of household, individual (belonging to the given household) and eligible women\textsuperscript{13} (of the household). The data set on eligible women provides variables on education of the married woman, education of husband, age, expectation of expenses borne on marriage by the groom and the bride side and other forms of gifts received from the community in marriage. These variables are used to assess assortative mating decisions on education and variations in dowry with female education. Factors such as household income and average education level in the household are mapped from the household data. These variables are used as controls since they may play an important role in selection of bride. For example, a household with a higher education on average would prefer an educated bride. Further religious affiliation and caste categories\textsuperscript{14} are also taken from the household data.

Marital decisions of an individual or a household are also governed by the socio-economic environment. Prominence of dowry and sex ratio (males per 1000 males) at the district level are used to control for this and taken from the Indian Census, 2011. Marriage expenditure net of expenses by the bride and groom’s family provides for a close proxy of dowry. Marital expenditure by the groom’s family was assessed by asking the head of the household, “At the time of the boy’s marriage, how much money is usually spent by the boy’s family?” Similarly, expenditure by the bride’s family was assessed by asking, “At the time of the girl’s marriage, how much money is usually spent by the girl’s family?” The responses were accounted as a range of minimum and maximum expenditure by each. Dowry was then calculated as the difference of average expenditure incurred by the bride and the groom’s family. To avoid measurement bias due to differences in cost of living across states, the calculated dowry was weighted by consumer price index of each state in 2011.

\textsuperscript{12} IHDS-I survey was conducted in 2005-2006. However, we use only IHDS-II since the other district variables are from Census of India, which is a decennial survey. These variables would have no variation across the two rounds of IHDS survey, 2005-06 and 2011-12.

\textsuperscript{13} Ever married women in the household in the age group of 15-49 years are surveyed in a face-to-face interview.

\textsuperscript{14} The Indian social system is characterised by the co-existence of multiple religious affiliations and hierarchy in castes with strong cultural and social beliefs that bear significant impact on socio-economic choices.
Marital expenditure however may not be a good measure for dowry prominence in a district. Marriage expenditure may be determined by other factors also such as ability to pay and preferences of individuals besides the norms of the society on marital transfers. Dowry prominence was therefore generated as a dummy from proportionate number of dowry deaths in 2011 in a district based on data provided by National Crime Records Bureau (NCRB). Dowry deaths provide a decent measure for the prominence of dowry in a district, since it is likely to be reported with a high probability and with greater accuracy. Amount of dowry paid is highly probable to be misreported and with a lesser likelihood. A district is defined as more intensive on dowry culture if the percentage of dowry deaths exceeds 1.5 (average percentage of dowry deaths)\textsuperscript{15}. Education levels of the husband and wife are measured as years of education completed. The questionnaire seeks information on “Years of Education completed” (of self and spouse).

The descriptive statistics for the variables from IHDS-II and Census 2011 are represented in Table 1.

\textbf{Table 1: Descriptive statistics for individual, household and district characteristics (IHDS-II)}

<table>
<thead>
<tr>
<th>Individual and Household</th>
<th>Obs.</th>
<th>Mean</th>
<th>Std. Dev.</th>
<th>Minimum</th>
<th>Maximum</th>
</tr>
</thead>
<tbody>
<tr>
<td>Wife Education</td>
<td>40349</td>
<td>5.14</td>
<td>4.90</td>
<td>0</td>
<td>16</td>
</tr>
<tr>
<td>Husband Education</td>
<td>37675</td>
<td>7.08</td>
<td>4.82</td>
<td>0</td>
<td>16</td>
</tr>
<tr>
<td>Wife’s Mother’s Education</td>
<td>40191</td>
<td>1.47</td>
<td>3.03</td>
<td>0</td>
<td>16</td>
</tr>
<tr>
<td>Husband’s Mother’s Education</td>
<td>40105</td>
<td>1.10</td>
<td>2.64</td>
<td>0</td>
<td>16</td>
</tr>
<tr>
<td>Average Education of Household</td>
<td>41912</td>
<td>7.04</td>
<td>4.55</td>
<td>0</td>
<td>16</td>
</tr>
<tr>
<td>Average Wedding Expenditure by Groom</td>
<td>45308</td>
<td>161.248</td>
<td>184,746</td>
<td>-</td>
<td>8,500,000</td>
</tr>
<tr>
<td>Average Wedding Expenditure by Bride</td>
<td>45302</td>
<td>248,641</td>
<td>257,202</td>
<td>-</td>
<td>8,500,000</td>
</tr>
<tr>
<td>Net Wedding Expenditure of Bride over Groom</td>
<td>45295</td>
<td>87,393</td>
<td>192,686</td>
<td>(6,750,000)</td>
<td>7,500,000</td>
</tr>
<tr>
<td>Household Income</td>
<td>45581</td>
<td>131,613</td>
<td>214,195</td>
<td>(1,037,040)</td>
<td>11,400,000</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>District</th>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>Number of Schools in District</td>
<td>186</td>
<td>352</td>
<td>273</td>
<td>30</td>
<td>1373</td>
</tr>
<tr>
<td>Sex Ratio (10-16 years)</td>
<td>365</td>
<td>1098</td>
<td>70</td>
<td>951</td>
<td>1331</td>
</tr>
<tr>
<td>Dowry Deaths (proportion of married females)</td>
<td>349</td>
<td>1.48</td>
<td>1.23</td>
<td>0</td>
<td>8.40</td>
</tr>
</tbody>
</table>

\textsuperscript{15} The variable requires due acknowledgement that dowry deaths may be due to weaker institutional arrangements, such as poor governance, weak criminal institutions, greater tolerance to violence against women and could influence the results. However, given the paucity of data, this is the best available proxy as a measure for dowry prominence in a district.
The descriptive statistics in Table 1 suggest that on an average, husbands are more educated than the wives. The net expenses borne on wedding by the bride’s family are positive, indicating positive groom price and marital transfers from the bride’s to the groom’s family. This may owe to relatively poor education levels and socio-economic status of the females in the country. On an average, the statistics are representative of sex-ratio biased towards males, showing 1098 males as compared to females in the age group of 10-16 years old. Statistics also indicate high dispersion in the number of high schools across districts. Dowry deaths as a percentage of married females in a district range from 0 to 8.4%. On an average, the districts in India witness approximately 1.5% of dowry deaths as a proportion of married girls.

Scatter diagram on the education level of husband and wife in Figure 1 shows a strong positive association between the two variables. Potential grooms and brides are likely to choose from their own education (ability) probability distribution.

Figure 1: Scatter plot for association between the education level of husband and wife in a household (IHDS-II)
Figure 2: Average dowry expenditure across states of India, IHDS-II

Figure 2 provides an overview of state-wise average net expenditure incurred by the bride’s family. The figure suggests large regional variation in the net marital expenditures borne by the bride side of the family. A closer look at the numbers is indicative of least dowry transfers in the north-eastern states of India such as Nagaland, Mizoram, Sikkim, Assam, etc. These states are also known for an above average level of women representation and independence. On the contrary, states from southern India such as Kerala, Tamil Nadu, Karnataka, etc., though possess highest standards of women literacy levels and women workforce participation, rank highest in dowry payments.

Region-wise and state-wise dot plot in Figure 3 provides an overview of the association between average dowry payments and years of schooling of the husband and wife. Dowry payments are seen to be higher in states with higher levels of husband education and lower levels of wife education. For example, Himachal Pradesh and NCT of Delhi in north both show similar husband years of schooling on an average. Data shows that dowry is however higher in Delhi with lower level of wife’s education. Similar inferences could be drawn for the states.
Figure 3: Region-wise dot plot of average dowry across states and average schooling years of husband and wife

4 Identification Strategy

To study the association between female education, marital sorting and dowry payments, the paper estimates the following specification for individual \( i \) in household \( h \) in state \( s \), district \( d \) belonging to religion \( r \) and caste \( c \)

\[
Education\ Wife_{ihd} = \alpha + \beta_1 Education\ Husband_{ihd} + \beta_2 Dowry\ District_d + \beta_3 Education\ Husband_{ihd}' Dowry\ District_d + \gamma_h X_h + \gamma_d X_d + \delta_{rh} + \delta_{ch} + \delta_s + \epsilon_{ihd} \quad (e1)
\]

where \( Education\ Husband_{ih} \) denotes the education level of husband (of the surveyed married female \( i \)) in household \( h \).
\( \beta_1 \) is a measure of assortative mating on education levels of husband and wife. A positive estimate of \( \beta_1 \) would indicate positive assortative mating where a household finds it ideal to marry potential grooms and brides with similar education levels. \( \beta_2 \) measures the difference in education level of brides in districts with greater prominence of dowry. Estimate of \( \beta_2 \), if positive, would indicate greater investment in female education in dowry prevalent districts in expectation of lower dowry payments at the time of marriage for a better educated bride. Estimate of \( \beta_2 \) may be negative, if female education expenditure is substituted for future dowry payment savings.

\( \beta_3 \) estimates the differential outcome on marital assortative decision of male education in districts with greater dowry prominence. Districts strongly driven by norms of dowry may exhibit a positive \( \beta_3 \) that would suggest stronger assortative mating in marriage market in dowry prevalent districts. In this case, dowry and female education are complements, such that higher dowry is paid to seek a better educated groom for a bride with higher education. If dowry and female education are substitutes, \( \beta_3 \) would be negative that indicates males with higher education find it desirable to substitute education of potential bride in compensation of dowry.

\( X_h \) and \( X_d \) are the household and district level controls such as income levels in the household, age gap of husband and wife, and urbanisation rate. \( \delta_{rh}, \delta_{ch} \) and \( \delta_s \) are the fixed effect controls for religion and caste of a household and state fixed effects, respectively.

Ordinary Least Square estimates may pose potential endogeneity bias due to omitted unobserved factors that may together determine the education levels of husband and wife. We thus derived instrumental variable estimates and instrument husband’s years of schooling with his mother’s years of schooling. Existing studies show that the education level of mothers is positively and significantly associated with the education level of children (Rihani, 2006; Coulon et. al., 2008). Husband’s mother’s years of schooling is derived from the “Eligible Woman” questionnaire for married females in the age group of 15-49 years. An eligible woman’s response to mother-in-law’s education is recorded from the question that asks, “How many standards/class your husband’s parents completed?” (The response is recorded for mother and father separately).

The analysis further directly tests for the relationship between dowry and female education in the districts of India. The objective is to examine if dowry and female education act as
substitutes or complements in marital decisions on education levels. $\beta_3$ coefficient in equation e2 directly tests for the association between female education and dowry.

$$\text{Log(Dowry)}_{hd} = \alpha + \beta_1 \text{Education Husband}_{ihd} + \beta_2 \text{Education Wife}_{ihd} + \beta_3 \text{Education Husband}_{ihd} \text{Education Wife}_{ihd} + \gamma_h X_h + \gamma_d X_d + \delta_{rh} + \delta_{ch} + \delta_s + \epsilon_{ih}$$ (e2)

$\beta_1$ is expected to be positive since higher education of potential groom would reflect greater ability and hence would demand larger dowry. $\beta_2$ on the contrary is expected to be negative. A rise in the education level of potential bride would raise her bargaining power in the marriage market and hence would lead to a fall in the transfers made to the groom’s family at the time of marriage. $\beta_3$ directly tests for the sign of $D_{mf}$, that is, how does dowry change with an additional year of education of both the groom and the bride. A positive estimate of $\beta_3$ would suggest complementarity between female education and dowry, such that the bride’s parents are willing to pay a higher dowry to seek a better ability husband for their daughters with higher education. If $\beta_3$ is negative, there exists substitutability between female education and dowry. With a positive $\beta_1$, though dowry may rise with male education, but it increases at a decreasing rate with a per unit rise in female education with a negative $\beta_3$. The household is willing to substitute dowry for a bride with higher education.

Unobserved factors may cause potential endogeneity in both female and male years of schooling. A female year of schooling is instrumented with the number of high schools in the district. Number of high schools in the district is expected to be positively correlated with female education years and exogenous as they are unlikely to be correlated with dowry exchanged in the household. The estimation further exploits the education of groom’s mother as an exogenous determinant of male education level\(^{16}\).

5 Empirical Assessment

Table 3A shows results for equation e1. The results are controlling for district and household specific characteristics that may impact the education level of the male in the family. Estimates on wife education are positive and highly significant which indicates positive

\(^{16}\) Bride’s mother’s years of schooling do not provide an exogenous determinant of her own education in a model that studies the association with dowry. The literature suggests that there exist strong inter-dependencies in a mother-daughter relation that may influence dowry exchanges through other factors besides daughter’s education levels and hence may not be entirely exogenous (Walker, Thompson & Morgan, 1987; Troll, 1987; Fischer, 1981).
assortative mating on education levels. Individuals tend to marry from their own ability distribution which suggests reinforcing effects on household cost reduction in the education of the bride and the groom.

**Table 3A: Marital assortative mating on years of schooling and difference in dowry prominent districts**

<table>
<thead>
<tr>
<th></th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>(4)</th>
<th>(5)</th>
<th>(6)</th>
<th>(7)</th>
<th>(8)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Wife’s Years of Schooling</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Husband’s Years of Schooling</td>
<td>0.606***</td>
<td>0.0935***</td>
<td>0.156***</td>
<td>1.301***</td>
<td>2.537***</td>
<td>2.653***</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.00407)</td>
<td>(0.00824)</td>
<td>(0.0128)</td>
<td>(0.0174)</td>
<td>(0.143)</td>
<td>(0.149)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>High Dowry Deaths Districts</td>
<td>0.108*</td>
<td>0.625***</td>
<td></td>
<td></td>
<td>0.0298</td>
<td>1.029**</td>
<td></td>
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</tr>
<tr>
<td></td>
<td>(0.0550)</td>
<td>(0.0953)</td>
<td></td>
<td></td>
<td>(0.161)</td>
<td>(0.435)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Dowry Death District*Husband’s Years of Schooling</td>
<td></td>
<td>-0.0702***</td>
<td></td>
<td></td>
<td>-0.144***</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.0108)</td>
<td></td>
<td></td>
<td>(0.0529)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Household Urban/Rural</td>
<td>0.739***</td>
<td>0.739***</td>
<td></td>
<td></td>
<td>0.566***</td>
<td>0.565***</td>
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<tr>
<td></td>
<td>(0.0398)</td>
<td>(0.0398)</td>
<td></td>
<td></td>
<td>(0.0975)</td>
<td>(0.0972)</td>
<td></td>
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<tr>
<td>Average Education of Household</td>
<td>0.730***</td>
<td>0.731***</td>
<td>-1.341***</td>
<td></td>
<td></td>
<td></td>
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<tr>
<td></td>
<td>(0.00952)</td>
<td>(0.00951)</td>
<td>(0.121)</td>
<td></td>
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<tr>
<td>Age Difference between Husband and Wife</td>
<td></td>
<td></td>
<td></td>
<td>-0.00819*</td>
<td></td>
<td></td>
<td>0.0192</td>
<td>0.0186</td>
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<td></td>
<td></td>
<td></td>
<td></td>
<td>(0.00466)</td>
<td></td>
<td></td>
<td>(0.0133)</td>
<td>(0.0133)</td>
</tr>
<tr>
<td>Log (Household Income)</td>
<td>-0.146***</td>
<td>-0.147***</td>
<td>-0.103**</td>
<td></td>
<td></td>
<td></td>
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<tr>
<td></td>
<td>(0.0194)</td>
<td>(0.0194)</td>
<td>(0.0489)</td>
<td></td>
<td></td>
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<tr>
<td>District FE</td>
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<td>Yes</td>
<td>No</td>
<td>Yes</td>
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<tr>
<td>State FE</td>
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<td>Yes</td>
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<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
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</tr>
<tr>
<td>Observations</td>
<td>37,612</td>
<td>35,985</td>
<td>35,985</td>
<td>37,415</td>
<td>35,795</td>
<td>35,795</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Robust standard errors in parentheses

*** p<0.01, ** p<0.05, * p<0.1

The results are in consensus with the findings of Mare (1991) and Mancuso (1997) who show that it is unlikely for individuals with higher education levels to marry someone with a lower education. Qian (1998) found that highly educated males were more likely to marry females with lesser education whereas females with higher education were less likely to marry someone with a lesser education level.

We derive interesting results from the study with the interaction variable between the dowry district and husband’s education. Though the estimates suggest a positive significant association between the education level of the husband and the wife, this association is weaker in the dowry prevalent districts as indicated by the negative significant effect of the interaction variable. The groom’s family is willing to accept a bride with lower education level for higher dowry. From the perspective of the girl’s family, there exist offsetting dowry effects in education since the parents of the bride, unable to secure a groom with higher
education in dowry prominent districts, substitute their daughter’s education for dowry. The results support King and Hill’s (1993) statement, “In cultures where dowry is customary, securing a more highly educated husband for an educated woman would require a larger dowry – another hidden cost of educating females”. If dowry and education were complements, the positive assortative mating results should have strengthened in districts where dowry is highly predominant.

Husband’s years of schooling may be endogenous as the wife’s level of education may determine the choice on husband’s level of education. This means that a female with higher education may choose to marry a male who is better educated and thus the causality may run in the reverse direction. Alternatively, husband and wife’s education levels may be determined simultaneously by factors such as cultural norms, preferences, attitude and perception of individuals towards education. Wu-Hausman’s F-statistic to test for the endogeneity of education level of husband is 685.43 and rejects the null of the variable being exogenous. The endogeneity is thus addressed using instrumental variable estimation as the identification strategy. Husband’s mother’s education is used as an IV for husband’s own years of schooling. The results are directionally consistent with OLS estimates, which suggest that there exists a positive association between the education level of the husband and the wife. However, OLS estimates are found to have a downward bias. It may be argued that husband’s mother’s education may not be entirely exogenous as the groom’s mother’s education level may alter her attitude toward the education level of her son’s future bride. However, statistics on a question to an eligible woman in the survey, “Please tell me who in your family decides: To whom your children should marry?” show that in 74% of the cases, the decision is made mostly by the “husband”. The respondent herself has a larger role only in 11% of the cases. Since females are less likely to make a decision on the marital choices for the children, it would be valid to assume that husband’s mother’s education is exogenous to the years of schooling of wife.

Table 3B: First Stage: Husband’s mother’s years of schooling as instrumental variable for husband’s years of schooling

<table>
<thead>
<tr>
<th></th>
<th>Husband’s Schooling</th>
<th>Husband’s Schooling*Dowry District</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
</tr>
<tr>
<td>Husband’s Mother’s Schooling</td>
<td>0.563***</td>
<td>0.101***</td>
</tr>
<tr>
<td></td>
<td>(.0087)</td>
<td>(.0066)</td>
</tr>
<tr>
<td>Dowry District Dummy</td>
<td>0.0723</td>
<td>0.0544</td>
</tr>
<tr>
<td></td>
<td>(.0631)</td>
<td>(.0668)</td>
</tr>
</tbody>
</table>
Table 3B shows the first stage estimates for instrumental variable estimation in Table 3A. First stage estimates for effect of husband’s mother’s years of schooling on husband’s schooling are shown in columns 1 to 3. The estimates suggest that men with mothers of higher education possess higher years of schooling. Husband’s mother’s education is significantly and positively associated with husband’s education level and the overall F-statistic of the model is 1804.21, indicating that the overall model is highly significant. Anderson-Rubin Wald test F-statistic is reported at 838.64. The test rejects the null hypothesis of weak identification and suggests that the husband’s mother’s education is a relevant instrument for husband’s own years of schooling. The instrument is statistically valid in the given model.

Table 4A: Association between dowry and years of schooling

<table>
<thead>
<tr>
<th></th>
<th>(1)</th>
<th>(2)</th>
<th>(1)</th>
<th>(2)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>OLS 1</td>
<td>OLS 2</td>
<td>IV 1</td>
<td>IV 2</td>
</tr>
<tr>
<td>Wife’s Years of Schooling</td>
<td>0.0362***</td>
<td>0.0295***</td>
<td>-0.0532**</td>
<td>-0.212***</td>
</tr>
<tr>
<td></td>
<td>(0.00130)</td>
<td>(0.00143)</td>
<td>(0.0235)</td>
<td>(0.0635)</td>
</tr>
<tr>
<td>Husband’s Years of Schooling</td>
<td>0.0209***</td>
<td>0.0164***</td>
<td>0.0804***</td>
<td>1.003***</td>
</tr>
<tr>
<td></td>
<td>(0.00233)</td>
<td>(0.00607)</td>
<td>(0.0171)</td>
<td>(0.349)</td>
</tr>
<tr>
<td>Wife’s Schooling*Husband’s Schooling</td>
<td>-0.000217</td>
<td>-0.0745**</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.000515)</td>
<td>(0.0302)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age difference Husband and Wife</td>
<td>0.00405**</td>
<td>-0.0124***</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.00168)</td>
<td>(0.00580)</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
Estimates in Table 4A directly test the association between dowry and education levels of husband and wife. OLS estimates in columns 1 and 2 show a positive significant relationship of education levels of both husband and wife with the amount of dowry exchanged. However, OLS estimates may be biased due to unobserved factors that may simultaneously affect the education level and dowry. Wu-Hausman F-statistic to test for exogeneity of wife’s years of schooling is 29.6 and rejects the null hypothesis and indicates the presence of endogenous factors. Instrumental variable estimates suggest that an additional year of husband’s schooling is associated with a 1 percent rise in dowry. A rise in wife’s schooling year on the contrary shows a fall in dowry by 0.21 percent. The estimates are consistent with the theory that postulates a rise in dowry with a rise in male education as a result of higher price for higher ability ($D_m > 0$). Dowry falls with female education as it raises the bargaining power of potential brides in the marriage market ($D_f < 0$). The coefficient for interaction between education levels of husband and wife tests for the joint marginal effect of additional years of schooling of the bride and groom on dowry. A negative estimate for the interaction term suggests that dowry is increasing at a decreasing rate. In reference to the theoretical model, this negative coefficient indicates that dowry and female education are substitutes such that $D_{mf} < 0$. These results provide credence to the assortative mating estimates in Table 3 that show that positive assortative mating result is weaker in dowry prominent districts. Since $D_{mf} < 0$, households tend to substitute female education with dowry, such that in districts where dowry is a strong social norm, females with lesser education get sorted with males of higher education. Dowry compensates for the lower education levels of the females.  

The analysis also aimed at directly testing joint effect of male and female education on the cost of household, $C_{mf}$. However, directly testing this parameter requires time use variables to measure effort and time in relation to education levels. The data for the same is not available in IHDS-II.
India therefore, transfers at the time of marriage are significant determinants of marital assortative mating. Other variables included in the analysis are directionally consistent. Those residing in urban areas and households with higher income show larger dowry payments. Large differences in the age of husband and wife amount to lesser dowry payments.

First Stage estimates are presented in Table 4B. The estimates show a strong positive association between the number of high schools in the district and wife’s years of schooling. Association with other variables is also consistent in reference to the existing literature. In case of higher district sex ratio with a greater number of males, females achieve lesser years of schooling. Estimates suggest that in households with urban residence and higher income, wives have higher education levels. Husband’s mother’s schooling level is also seen to be positively associated with the wives’ schooling years. Higher age difference between husband and wife is negatively correlated to wives’ education level.

<table>
<thead>
<tr>
<th></th>
<th>Wife’s Schooling</th>
<th>Wife’s Schooling*Husband’s Mother’s Schooling</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>IV 1</td>
<td>IV 2</td>
</tr>
<tr>
<td>Number of High Schools in District</td>
<td>.0018***</td>
<td>.0013***</td>
</tr>
<tr>
<td></td>
<td>(.0001)</td>
<td>(.00019)</td>
</tr>
<tr>
<td>Husband’s Mother’s Schooling</td>
<td>.7115***</td>
<td>4.2216***</td>
</tr>
<tr>
<td></td>
<td>(.0135)</td>
<td>(.16768)</td>
</tr>
<tr>
<td>Age difference between Husband and Wife</td>
<td>-.0461***</td>
<td>-.0084**</td>
</tr>
<tr>
<td></td>
<td>(.01056)</td>
<td>(.00427)</td>
</tr>
<tr>
<td>Log (Household Income)</td>
<td>.7598***</td>
<td>.18435***</td>
</tr>
<tr>
<td></td>
<td>(.0376)</td>
<td>(.0152)</td>
</tr>
<tr>
<td>District Sex Ratio</td>
<td>-.0069***</td>
<td>-.0006</td>
</tr>
<tr>
<td></td>
<td>(.00108)</td>
<td>(.00043)</td>
</tr>
<tr>
<td>Household Residence: Urban</td>
<td>1.7590***</td>
<td>.3296***</td>
</tr>
<tr>
<td></td>
<td>(.07920)</td>
<td>(.0320)</td>
</tr>
<tr>
<td>Number of High Schools*Husband’s Mother’s Schooling</td>
<td>-.00071***</td>
<td>.00054***</td>
</tr>
<tr>
<td></td>
<td>(.00032)</td>
<td>(.00013)</td>
</tr>
</tbody>
</table>

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

Cragg-Donald F-statistic is reported at 12.32 which rejects the hypothesis of weak identification and suggests that the number of high schools is a relevant instrument for wife’s years of schooling. The overall F-statistic of first stage model is 210.92 and indicates the overall model being significant.
Robustness Tests

It may be argued that the results on positive assortative mating may be driven by females on the higher end of ability distribution, with higher years of schooling and in the workforce. In other words, marital assortative mating is driven by greater expectation of future wage income of wives as suggested by literature. We derive estimates for assortative mating on education by excluding females ‘who have ever worked for pay/wages’ and with years of schooling more than 10 years. The estimates in Table 5 provide association between years of schooling for husband and wife for a sub-sample where females either do not have sufficient education levels or have never generated any income, such that female education offers pure non-monetary benefits to the household. The estimates are consistent with the main specification, which suggests positive marital assortative mating in education, with results being weaker in districts with greater dowry prominence.

Table 5: Robustness validation instrumental variable estimates for marital assortative mating based on husband and wife’s years of schooling

<table>
<thead>
<tr>
<th>Wife’s Years of Schooling</th>
<th>Excludes women who have ever worked for wages</th>
<th>Excludes women with years of schooling&gt;10</th>
</tr>
</thead>
<tbody>
<tr>
<td>Husband’s Years of Schooling</td>
<td>1.352*** (0.0201)</td>
<td>1.504*** (0.0557)</td>
</tr>
<tr>
<td>Dowry Death District*Husband’s Years of Schooling</td>
<td>-0.211*** (0.0552)</td>
<td>-0.179** (0.0750)</td>
</tr>
<tr>
<td>High Dowry Deaths Districts</td>
<td>1.608*** (0.494)</td>
<td>1.089** (0.441)</td>
</tr>
<tr>
<td>District FE</td>
<td>Yes</td>
<td>No</td>
</tr>
<tr>
<td>State FE</td>
<td>No</td>
<td>Yes</td>
</tr>
<tr>
<td>Includes Controls</td>
<td>No</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>22,239</td>
<td>21,932</td>
</tr>
</tbody>
</table>

Robust standard errors in parentheses

*** p<0.01, ** p<0.05, * p<0.1

It may be argued that the instrumental variable estimates derived in Table 4A, that directly tests for the joint association of husband and wife’s years of schooling on dowry, may be driven by certain factors specific to individuals, household or the districts themselves. Robustness validation tests are presented in Table 6 to validate the robustness of estimates from the main model. Mother-in-law’s schooling years are taken as an exogenous determinant of husband’s education. Since dowry is measured as a perception of the head of the household, mother-in-law’s education might not be completely exogenous if she herself is the head of the household. Models 1 and 2 therefore exclude households with female heads to
remove simultaneity between mother-in-law’s education and dowry. The estimates are found to be significant and close to the coefficients in the main model.

Models 3 and 4 exclude women from analysis who belonged to a different town or city before marriage so that the instrument, i.e., ‘number of high schools in the district’ for wife’s education level is relevant. The IHDS-II data does not provide information on the district of residence of females before marriage. If a large number of women belonged to a different district, the IV would be irrelevant. Including only those females in the analysis that belong to the same district before marriage ensures that the IV is relevant for the eligible women’s education levels.

Table 6: Robustness validation instrumental variable estimates for the association between dowry payments and years of schooling

<table>
<thead>
<tr>
<th></th>
<th>Exclude Household with Female Heads</th>
<th>Exclude Women from different Districts before marriage</th>
<th>Exclude Outlier Dowry Deaths Districts</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
<td>(3)</td>
</tr>
<tr>
<td>Wife’s Years of Schooling</td>
<td>-0.0632*** (0.0240)</td>
<td>-0.216*** (0.0643)</td>
<td>-0.0721*** (0.0266)</td>
</tr>
<tr>
<td>Wife’s Schooling*Husband’s Schooling</td>
<td>-0.0775** (0.0317)</td>
<td>-0.0700*** (0.0271)</td>
<td>-0.0726** (0.0354)</td>
</tr>
<tr>
<td>Husband’s Years of Schooling</td>
<td>0.0878*** (0.0174)</td>
<td>1.041*** (0.366)</td>
<td>0.0929*** (0.0194)</td>
</tr>
<tr>
<td>State FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Include Controls</td>
<td>No</td>
<td>Yes</td>
<td>No</td>
</tr>
<tr>
<td>Observations</td>
<td>14,819</td>
<td>14,342</td>
<td>13,033</td>
</tr>
</tbody>
</table>

Robust Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

Models 5 and 6 exclude the outlier districts in dowry deaths to ensure that the results are not driven by extreme values. Excluding the districts that either have a history of almost negligible dowry deaths or a large percentage of dowry deaths, we exclude any biasedness in the results that may be caused due to strong cultural norms or social beliefs.

The estimates from the validation tests are consistent, significant and similar to the main model. It validates that dowry payments are significantly and positively associated to male education. However, the rate of increase is lesser with additional years of female education.
6 Conclusion

Assortative mating in marriage market is a function of characteristics of potential brides and grooms and their respective households. The aim is to maximise household utility which is a function of both monetary and non-monetary traits. Monetary traits such as higher wages raise income and positively impact utility, whereas non-monetary traits such as appearance, intelligence, education, etc. bring down the household cost of production, thereby raising net utility. As proposed by Anderson (2007), households gain higher utility from brides and grooms with higher ability. Existing literature provides evidence of positive assortative mating on education. The studies argue that men with higher education prefer educated females in expectation of labour market wage income. However, in developing countries such as India factors such as poor female workforce participation, strong parental roles in marriage related decisions and existence of dowry as a social norm play a dominant role in marital choices relative to potential wage income of the bride. Female education is perceived as a non-monetary trait that leads to positive marginal utility in household production and fall in the cost of production. Therefore, on one hand, education of females would raise the utility of the post-marital household. On the other hand, there may exist substitutability between the education level of the bride and dowry transfers. A rise in female education may lead to a fall in dowry receipts by the groom’s family. In theory, dowry acts as an equilibrating mechanism in the marriage market and can compensate for the differences in the ability gaps of the potential grooms and brides. Though dowry is strongly driven by social and cultural norms, yet the magnitude varies with the abilities of the potential groom and bride. The higher is the ability level of the groom, the larger is the net marital transfer from the bride to the groom’s side. Thus, there exists a strong relationship between education levels of potential bride and marital assortative mating. The outcome may vary with the strength of substitution effect between female education and dowry payments.

An empirical assessment of assortative mating and dowry with rise in the education level of females in India using IHDS-II data suggests significant positive assortative mating on female education. Estimates suggest that brides and grooms tend to choose partners from their own ability distribution. This association however is found to be weaker in districts with relatively higher percentage of dowry deaths. This paper proposes a basic theoretical model to support the empirical estimates. We suggest that a rise in female education would impact household utility or production through both a fall in dowry and reduction in cost of household production. On one hand, fall in dowry reduces the post-marital family income and
hence utility and, on the other hand, a reduction in household cost of production would lead to an increase in household production, per unit of cost. A household should always gain from a rise in the education level of a male through both rise in dowry and fall in production cost, however, bride’s education would have countering effects. Marital assortative mating decisions based on the education of potential bride and groom together would be determined by the cross marginal effects of education on household production. Magnitude and direction of cross effects of male and female education on dowry and household costs would determine the net effect on assortative mating. A rise in education of male and female, together, may either lead to offsetting or reinforcing effects on cost and dowry. Effects would be reinforcing in cost if a rise in education of both the bride and the groom generates increasing returns to scale and cost falls at a faster rate. The cost effects would be offsetting in the presence of decreasing returns to scale. Dowry may rise with male education but at a lower rate if the female is more educated. This generates offsetting effects on dowry. Parents of the bride on the other hand may be willing to pay a higher dowry for a better educated husband as the daughter attains higher education. In this case, dowry effect in education is reinforcing. The model suggests that positive assortative mating would be desirable for a household if the efficiencies achieved in cost are greater than the loss in dowry with rise in female education.

The empirical results for India using IHDS-II data provide evidence of positive assortative mating and suggest that the cost effects outweigh the fall in dowry effect. The effect is however found to be weaker in dowry prominent districts. This indicates that as social and cultural norms of dowry strengthen, substitutability between dowry and female education also increases. The marriage market allows dowry to compensate for relatively low education level of the bride. The analysis directly tests for joint association between male and female education and dowry. The estimates show that the dowry payments rise with additional years of male schooling and fall with female education. The effect of male and female education interaction variable on dowry is negative, indicating that dowry increases with male education but at a decreasing rate in response to rise in female education. Households are willing to accept lower dowry payments for a better educated bride. Probable endogeneity bias is addressed through instrumental variable estimation where the ‘number of high schools’ in a district is used as instrumental variable for female education. The estimates are found to be robust to validation tests which exclude factors specific to individuals, households or districts that may have driven the results.
References


Gihleb, R., & Lang, K. (2016). Educational homogamy and assortative mating have not increased (No. w22927). National Bureau of Economic Research.


**Appendix**

A.1

**Derivation of Equation (7)**

\[
\frac{\partial Z}{\partial s_m} = \frac{\partial [D(s_m, s_f)C^{-1}]}{\partial s_m}
\]
\[
\frac{\partial D}{\partial S_m} c^{-1} - c^{-2} \frac{\partial C}{\partial S_m} D(S_m, S_f)
\]

Household production would increase with male years of schooling if:

\[D_m - C^{-1}C_m D(S_m, S_f) > 0\]

Dividing both sides by \(D\) and multiplying by \(S_m\)

\[\frac{\partial D}{\partial S_m} S_m D - \frac{\partial C}{\partial S_m} S_m C > 0\]

\[\tau_m - \omega_m > 0\]

Where \(\tau_m = \frac{\partial D}{\partial S_m} \cdot \frac{S_m}{D}\) and \(\omega_m = \frac{\partial C}{\partial S_m} \cdot \frac{S_m}{C}\)

**Derivation of Equation (8)**

\[
\frac{\partial Z}{\partial S_f} = \frac{\partial [D(S_m, S_f)c^{-1}]}{\partial S_f}
\]

\[= \frac{\partial D}{\partial S_f} c^{-1} - c^{-2} \frac{\partial C}{\partial S_f} D(S_m, S_f)\]

Household production would increase with female years of schooling if:

\[D_f - C^{-1}C_f D(S_m, S_f) > 0\]

Dividing both sides by \(D\) and multiplying by \(S_f\)

\[\frac{\partial D}{\partial S_f} S_f D - \frac{\partial C}{\partial S_f} S_f C > 0\]

\[\tau_f - \omega_f > 0\]

Where \(\tau_f = \frac{\partial D}{\partial S_f} \cdot \frac{S_f}{D}\) and \(\omega_f = \frac{\partial C}{\partial S_f} \cdot \frac{S_f}{C}\)

**A.2**

**Derivation of Equation (10)**
\[
\frac{\partial Z}{\partial S_m} = \frac{\partial}{\partial S_m} [D(S_m,S_f)C^{-1}]
\]
\[
= \frac{\partial D}{\partial S_m} C^{-1} - C^{-2} \frac{\partial C}{\partial S_m} D(S_m,S_f)
\]
\[
= D_m C^{-1} - C_m^{-2} C D(S_m,S_f)
\]

\[
\frac{\partial^2 Z}{\partial S_m S_f} = D_{m,f} C^{-1} - C^{-2} D_m C_f - [C^{-2} C_m D_f - 2 C^{-3} C_m C_f D + C^{-2} C_{m,f} D]
\]
\[
= C^{-1} D_{m,f} - C^{-2} D_m C_f - C^{-2} C_m D_f + 2 C^{-3} C_m C_f D - C^{-2} C_{m,f} D
\]

A.3

**Derivation of Equation (11')**

\[
2 C^{-2} C_m C_f D - C^{-1} D_m C_f > C^{-1} C_m D + C^{-1} C_m D_f - D_{m,f}
\] (11)

Multiplying both sides by \( S_m \) and dividing by \( D \)

\[
2 \frac{\partial C}{\partial S_m} \frac{S_m}{C} = \frac{\partial D}{\partial S_m} \frac{S_m}{D} > C \left[ \frac{C_{m,f} D S_m}{C. D. C_f} + \frac{C_m D_f S_m}{C. D. C_f} - \frac{D_{m,f} S_m}{D. C_f} \right]
\]

The left hand side could be written as:

\[
2 \omega_m - \tau_m
\]

We have three terms on the right sides of the equation.

The first term on the right hand side could be written as:

\[
C \left[ \frac{C_{m,f} D S_m}{C. D. C_f} \right]
\]

\[
= \frac{\partial}{\partial S_m} \left( \frac{\partial C}{\partial S_f} \right) S_m
\]

\[
= \frac{\partial C_f}{\partial S_m} \frac{S_m}{C_f}
\]

\[
= \omega_{f,m}
\]
The second term on the right hand side could be written as:

\[ C \left[ \frac{C_m D_f S_m}{C \cdot D \cdot C_f} \right] \]

Multiplying and dividing by \( S_f \)

\[
= \frac{C \frac{\partial c}{\partial s_m} \frac{\partial D}{\partial s_f} S_m}{C \cdot D \frac{\partial c}{\partial s_f} S_f}
\]

\[ = \frac{\omega_m \cdot \tau_f}{\omega_f} \]

Third term on the right hand side could be written as:

\[ C \left[ \frac{D_{m,f} S_m}{D \cdot C_f} \right] \]

\[
= \frac{\frac{\partial (\frac{\partial D}{\partial s_m})}{\partial s_f} S_m}{D \cdot C_f}
\]

\[ = \frac{\frac{\partial D_m}{\partial s_f} S_m}{D \frac{\partial c}{c \cdot \partial s_f}} \]

In the numerator, multiplying and dividing by \( D_m \) and \( S_f \)

\[
= \frac{\frac{\partial D_m}{\partial s_f} \frac{D_m}{D_{m,f} S_f} \frac{S_m}{S_f}}{D \frac{\partial c}{c \cdot \partial s_f}}
\]

\[ = \frac{\tau_{m,f} \cdot \tau_m}{\omega_f} \]

Equation (11) therefore could be written as:

\[
2\omega_m \omega_f - \tau_m \omega_f > \omega_{f,m} \omega_f + \omega_m \cdot \tau_f - \tau_{m,f} \cdot \tau_m \quad (11')
\]